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Bernice C. Yates

Carol H. Pullen

Jonathan Santo

Linda Boeckner

Patricia A. Hageman

See next page for additional authors

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Authors

Bernice C. Yates, Carol H. Pullen, Jonathan Santo, Linda Boeckner, Patricia A. Hageman, Paul J. Dizona, and Susan Noble Walker

The influence of cognitive-perceptual variables on patterns of change over time in rural midlife and older women's healthy eating

Bernice C. Yates a,*, Carol H. Pullen a, Jonathan Bruce Santo b, Linda Boeckner c, Patricia A. Hageman d, Paul J. Dizona a, Susan Noble Walker a

- ^a College of Nursing, University of Nebraska, 985330 Nebraska Medical Center, Omaha, NE 68198-5330, USA
- b Department of Psychology, University of Nebraska at Omaha, USA
- ^c University of Nebraska Panhandle Research and Extension Center, Scottsbluff, USA
- ^d Department of Allied Health, University of Nebraska Medical Center, Omaha, USA
- * Corresponding author. Tel.: þ1 402 559 5460; fax: þ1 402 559 8188.

E-mail address: bcyates@unmc.edu (B.C. Yates).

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Abstract

Although studies demonstrate that dietary interventions for healthy adults can result in beneficial dietary changes, few studies examine when and how people change in response to these interventions, particularly in rural populations. The purpose of this study was to examine patterns of change over time in healthy eating behaviors in midlife and older women in response to a one-year health-promoting intervention, and to examine what predictors (perceived benefits, barriers, self-efficacy, and family support for healthy eating) influence the changes during the intervention and follow-up. Data for this secondary analysis were from the Wellness for Women community-based trial. Women (N = 225) between the ages of 50–69 in rural Nebraska, U.S.A., were recruited. A repeated-measures experimental design was used with randomization of two rural counties to intervention (tailored newsletter) or comparison (standard newsletter) groups. Eating behavior was measured by the Healthy Eating Index. The predictor variables were assessed using standard measures. Data analysis was done using latent growth curve modeling. The tailored

newsletter group was successful in improving their healthy eating behavior compared to the standard newsletter group during the one-year intervention, at the end of the intervention, and during the follow-up phase. Family support at the end of the intervention was positively associated with healthy eating at the end of the intervention. Perceived barriers had the strongest impact on healthy eating behavior at all time points. Compared to participants in the standard newsletter group, those in the tailored newsletter group perceived more family support and fewer barriers for healthy eating at the end of the intervention (mediation effects). Based on these findings, both family support and perceived barriers should be central components of interventions focused on healthy eating behavior in rural midlife and older women.

Introduction

Health-promoting interventions are critically important to improve healthy eating and prevent disease in individuals. Studies demonstrate that dietary interventions for healthy adults can result in beneficial dietary changes; however, few studies examined when and how people change in response to these interventions (Brunner, Rees, Ward, Burke, & Thorogood, 2007; Lin, O'Connor, Whitlock, & Beil, 2010). Fewer studies have examined factors influencing dietary change in rural populations. Thus, the purpose of this secondary analysis was to examine patterns of change over time in healthy eating habits in rural midlife and older women in response to a one- year health-promoting intervention, and to examine what determinants influence those changes during the intervention and follow-up. The cultural context of this investigation was the rural environment of Nebraska, U.S.A., where the Wellness for Women community-based study was conducted. Of the 93 counties in Nebraska, 85 are considered rural, including 32 counties that are sparsely populated frontier areas. Rural women report poorer health status (Feresu, Zhang, Puumala, Ullrich, & Anderson, 2008), higher rates of obesity (Henderson & Low, 2006), consume higher fat diets (Boeckner, Pullen, Walker, Oberdorfer, & Hageman, 2007), and have lower rates of access to and use of health care services (McCall-Hosenfeld & Weisman, 2011).

The theoretical approach for this study was a Health Promotion Model (HPM) (Pender, Murdaugh, & Parsons, 2006) that included four cognitive-perceptual factors as determinants of healthy eating behavior: benefits, barriers, self-efficacy, and family support. All four factors were found to influence dietary behaviors in descriptive studies. Women who perceived greater benefits, self-efficacy, family support and fewer barriers for healthy eating were more likely to consume fruits, vegetables, whole grains, and to limit intake of meat and fats (Walker, Pullen, Hertzog, Boeckner, & Hageman, 2006). Pawlak and Colby (2009) found that perceived benefits of healthy eating generally received higher scores than perceived barriers to healthy eating in African Americans. Kaiser, Brown, and Baumann (2010) found that rural, low-income individuals who perceived low barriers to healthy eating were significantly more likely to meet the recommended intake for fruits and vegetables. In 4 rural communities in Southwest Georgia, women with greater self-efficacy to eat healthy foods reported less high fat food in the home and lower levels of dietary fat intake (Hermstad, Swan, Kegler, Barnette, & Glanz, 2010). Similarly, family support was associated with increased fruit and vegetable consumption (Luszczynska & Cieslak, 2009) and reduced fat intake (Hagler et al., 2007).

Similar to findings from descriptive studies, perceived benefits (Kelley & Abraham, 2004; Resnicow et al., 2000), barriers (Henson, Blandon, & Cranfield, 2010), self-efficacy (Folta et al., 2009; Mosher et al., 2008) and support for healthy eating (Barrera, Strycker, MacKinnon, & Toobert, 2008) have all been shown to influence dietary behaviors in intervention studies. Noar, Benac, and Harris (2007) conducted a meta-analysis of tailored interventions focused on a wide variety of health behaviors such as smoking cessation, dietary change, and mammography screening. They found that the average effect size (ES) of tailoring on health behavior change was r = .074 (somewhat less than Cohen's small ES of r = .10) suggesting that tailored messages had a greater impact on health behavior than comparison groups. The effect size for tailored interventions on healthy eating behavior was r = .084 (Noar et al.). Of the 57 studies reviewed, 44 used a single intervention contact only. In the remaining 13 studies, the number of contacts ranged from 2 to 12 (median = 3). The most effective interventions were those with >1 contact, used pamphlets or newsletters rather than manuals or

booklets, and tailored on 4-5 concepts rather than 0-3. Concepts associated with larger effect sizes included attitudes, stages and processes of change, self-efficacy and social support (Noar) of which the latter two are included in this study.

In a recent meta-analysis of the effects of behavioral counseling to promote a healthy diet to prevent cardiovascular disease in adults, Lin et al. (2010) found that interventions significant beyond the first few months in reducing fat and saturated intake and/or increasing fruit and vegetable intake were all high-intensity counseling interventions (>360 min) with group, phone, or mail contact throughout the trial. Of the 73 studies reviewed, however, only 11 studies followed outcomes beyond 12 months (Lin). In the Wellness for Women study, Walker et al. (2009, 2010) found that women in the tailored newsletter group, where messages were based on women's individual assessments, made greater improvements in healthy eating (more fruits and vegetables, reduced fat intake) during the 12 month intervention than women in the standard newsletter group (Walker et al., 2009). From 12 to 24 months, however, both groups maintained levels of healthy eating and no group differences were found (Walker et al., 2010). Potential group differences may have been masked by examining separate food categories (intake of fat, fiber, fruits and vegetables) rather than diet as a whole. Thus, in this study, a composite outcome measure of dietary quality was used.

In their seminal work with rural populations, Long and Weinert (1989) found that the culture of rural dwellers centered on the key concepts of work and health beliefs, self-reliance and independence, and isolation and distance. These culturally rooted beliefs and living situations may negatively influence health behaviors. McCall-Hosenfeld and Weisman (2011) compared use of preventive counseling in rural vs. urban women and found that rural women (32%) were significantly less likely to receive nutrition counseling compared with urban women (40%). Another study found that rural dwellers do not use preventive services to the same extent as urban dwellers (Pullen, Fiandt, & Walker, 2001). The culture of rural isolation may contribute to altered eating patterns that can challenge women's development of healthy lifestyle patterns (Bove & Olson, 2006). In many rural areas, the availability of fresh fruits and vegetables is limited in the winter months and grocery stores with these items may be at a distance.

Nebraska Behavioral Risk Factor Surveillance System data indicate that only 25.8% of women eat 2:5 fruits and vegetables daily. Further, fruit and vegetable consumption decreases with increasing age in Nebraska adults (Nebraska Office of Women's and Men's Health Report, 2009-2010).

In summary, although there is consensus that benefits, barriers, self-efficacy, and family support serve as major determinants of healthy lifestyle behavior, there is little consensus about the causal paths linking these variables to behavior. In this study, we want to empirically identify the variables that most strongly influence actual performance of healthy eating behavior. There is a need to establish when and how people change in response to an intervention to focus efforts during the most appropriate time. While theoretical models that describe the determinants and processes of health behavior change are readily available, analyzing these models using latent growth curve modeling will provide insight into how various determinants bring about change in health behavior.

The purpose of this study was to examine patterns of change over time in healthy eating habits in response to a one-year intervention in two treatment groups (tailored and standard newsletter groups); and to examine how determinants influence those changes at each phase of the intervention and follow-up. The specific research questions were:

1) What are the patterns of change during the intervention (baseline, 3, 6, 9, and 12 months) and maintenance period (18 and 24 months) in healthy eating behaviors in response to a one-year intervention in both tailored and standard newsletter groups? 2) How do the significant determinants (benefits, barriers, self- efficacy, and family support for healthy eating) change during the intervention, and how do they influence changes in healthy eating behavior during and at the end of the intervention and the maintenance period? 3) Which determinants are significant mediators between the tailored vs. standard newsletter groups and healthy eating behavior during and at the end of the intervention and the maintenance period?

Methods

The Wellness for Women study was a community-based study using a

repeated-measures experimental comparison group design. Women were recruited from two similar rural counties in separate corners of Nebraska. The geographic areas rather than participants were randomized to intervention (tailored newsletters) or comparison (standard newsletter) groups to avoid cross- intervention contamination in these rural areas where women tend to frequent a limited number of commercial and social settings. Although the initial study examined physical activity and eating behaviors; only the eating behaviors are reported here.

Sample

Two hundred and twenty five women were recruited for the study. Inclusion criteria were: a) women aged 50-69; b) English speaking; and c) able to utilize a computer at rural research offices to complete online surveys. To ensure that all participants would have the potential for change in eating behaviors, women were excluded if they were in the maintenance stage for consumption of fat, fruits and vegetables, and grain intake as measured by the Stages of Healthy Eating Questionnaire (adapted from Marcus, Rossi, Selby, Niaura, & Abrams, 1992). More details about participant recruitment can be found in Walker et al. (2009). The study was approved by the University's Institutional Review Board; each participant provided written informed consent. Although measures included in this study were self-report, all measures were taken by RNs during visits to the rural research office, during the intervention (0, 3, 6, 9 and 12 months) and every 6 months during follow-up (18 and 24 months), at which time biomarkers were also assessed. The RNs were blinded to group assignment and were instructed to clarify items on the self-report measures if necessary. They were also specifically instructed to refrain from any type of counseling.

Sample size

In the original study, power analyses were done for the primary outcomes of changes in eating and activity behavior. Using repeated-measures analysis of variance (RMANOVA), small-to-medium differences were detected in the amount of change between groups (partial n^2 effect size = .04), alpha = .05, and power \geq .80 (Walker et al., 2009). A larger sample size is typically needed for RMANOVA than for latent

growth curve (LGC) analysis because LGC modeling accounts for within subject variability better than RMANOVA; thus, fewer participants are needed (Little, in press).

Theory-based intervention

Four behavior-specific cognitions from the HPM shaped the intervention strategies for this study: benefits, barriers, self- efficacy, and interpersonal influences for healthy eating (Bandura, 1997; Pender et al., 2006). These cognitive factors influence behavior directly and are considered modifiable, an essential characteristic of determinants proposed as a basis for structuring interventions to promote health behavior change. The intervention lasted 12 months with an additional 12 month followup. Eighteen mailed tailored or standard newsletters were sent to women every 2 weeks for the first 6 months and every 4 weeks for the next 6 months. Both newsletters were written with similar layout, font, and length to minimize differences in appearance. The standard newsletters contained general information about healthy eating that is currently available from organizations such as the American Heart Association and American Cancer Society. For the tailored newsletters, a library of hundreds of messages was created that corresponded to individual women's data. The content of the tailored newsletters was individualized in relation to: Personal goals, most current assessment of benefits, barriers, self-efficacy, and interpersonal support, and biomarker results regarding eating. Of the 18 newsletters, the benefits of healthy eating were discussed in 7, barriers were discussed in 11, self-efficacy was dis- cussed in 16, and interpersonal support was discussed in 6. The daily dietary components emphasized in the intervention were: 2 servings of fruits; 3 servings of vegetables (1 dark green or deep yellow); 6 servings of grains (3 whole grains); ≤ 30% of calories from fat; and <10% of calories from saturated fat.

Measurement of predictor variables

For the predictor variables in the measurement model, items were selected specifically because they were theoretically relevant and salient to the concept of interest. We then examined their reliability scores using Cronbach's alpha, how well the items factored together, and selected the ones that were the most

statistically similar. Once the indicators at each time point were found to be reliable and valid using confirmatory factor analysis, a composite score was created and used as the observed variable for the models.

Perceived benefits and barriers to healthy eating were measured by selected items from the Healthy Eating Benefits and Barriers Scales (HEBBS) (Sechrist, Walker, & Pender, 1987). These scales used a 4-point response ranging from 1 (strongly disagree) to 4 (strongly agree). Construct validity, internal consistency, and test-retest reliability were reported by Sechrist et al. (1987). In this study, benefits were measured by 4 items: Healthy eating helps me to be more fit, helps me to have more energy, helps me to lose weight, and makes me feel better. Barriers were measured by 5 items: Healthy eating is inconvenient, too expensive, takes too much time to shop for healthy foods, takes too much time to prepare healthy meals, and means giving up foods that I like. Cronbach's alphas for the benefits items ranged from .76 to .84 and for the barriers items ranged from .74 to .82.

Self-efficacy for healthy eating was measured by selected items from the Self-Efficacy for Eating Habits Scale (Sallis, Pinski, Grossman, Patterson, & Nader, 1988). This scale measures individuals' beliefs in their confidence to perform healthy eating behavior under various circumstances. The 6-items were: How confident are you that you would avoid eating fast food for lunch or dinner, eat carrots, celery and raw vegetables instead of dips, crackers, and potato chips for snacks, avoid junk food that other family members have brought into your home, stick to low fat foods when high fat food is readily available at a party, stick to low fat foods when you are alone and there is no one there to watch you, and stick to low fat foods when you feel depressed, bored or tense? Self-efficacy was rated on a 5-point Likert scale ranging from 1 (I know I cannot) to 5 (I know I can). Evidence of construct validity, test-retest reliability, and predictive criterion-related validity were reported (Calfas, Sallis, Oldenburg, & Ffrench, 1997; Sallis et al., 1988). In this study, Cronbach's alphas ranged from .81 to .88.

Family support of healthy eating was measured by the Family Support for Healthy Eating Habits Scale (Calfas et al., 1997; Sallis, Grossman, Pinski, Patterson, & Nader, 1987). This scale measures positive encouragement for healthy eating provided by family members. In this analysis, family support for healthy eating was measured by

4 items: My family commented if I went back to my old eating habits, discussed my eating habit changes with me, reminded me not to eat high fat foods, and encouraged me not to eat "unhealthy foods" (for example, cake or fried foods) when I was tempted to do so. Items were measured on a 5-point Likert scale ranging from 1 (none) to 5 (very often) to indicate the frequency of support provided. Higher scores indicated greater family support for healthy eating. The scale has evidence of construct validity, test-retest reliability, and predictive criterion-related validity (Calfas et al., 1997; Sallis et al., 1987). In this study, Cronbach's alphas ranged from .82 to .86.

Outcome: healthy eating index

The web-based version of the 1998 Block Health Habits and History Questionnaire was used to measure eating behavior (Boeckner, Pullen, Walker, Abbot, & Block, 2002). This food frequency questionnaire (FFQ) provides estimates of the nutrients and dietary constituents in a person's diet. The anchors for the FFQ assessments were: the past year for baseline, past 3 months during the intervention phase, and past 6 months during the maintenance phase. Rather than examine a single food constituent (i.e., fruits and vegetables), we used the Healthy Eating Index (HEI) 2000 (Basiotis, Carlson, Gerrior, Juan, & Lino, 2002) to create a composite score of dietary quality. Typically diet quality indices like the HEI have been calculated from food recalls and records. However, several investigators have successfully calculated and validated HEI scores from FFQ's (Boynton et al., 2007; Feskanich, Rockett, & Colditz, 2004; Savoca et al., 2009; Shatenstein, Nadon, Godin, & Ferland, 2005) similar to the method used in this study. The HEI measures how well American diets conform to recommended healthy eating patterns and is designed to measure overall dietary quality. Each of the 10 components (grains, vegetables, fruits, milk, meat, total fat, saturated fat, cholesterol, sodium, and variety) has a scoring range of 0e10. Scores for each component were assigned based on the recommended number of food guide pyramid servings per day for women ≥51 years of age (e.g., a maximum score of 10 points was given for grains = 7.4 servings, vegetables = 3.5, fruits = 2.5, milk = 2, and meat = 2.2) (Achterberg, McDonnell, & Bagby, 1994; Basiotis et al., 2002). In addition, a score of 10 was given for intakes of fat \leq 30%, saturated fat \leq 10%,

cholesterol ≤300 mg, and sodium ≤2400 mg. A score of zero was assigned when no foods in a particular group were eaten (or when the following levels were achieved: 45% fat, 15% saturated fat, 450 mg cholesterol, and 4800 mg of sodium) (Basiotis et al., 2002). Intermediate scores were calculated proportionately. Variety was calculated by taking 5 food groups (grains, fruits, vegetables, dairy, and meats), counting the number of different foods they had at least a half serving of in a day, and coding on the 0 to 10 scale. The overall HEI has a total possible score ranging from 0 to 100, with 100 indicating best dietary quality. Validity evidence was provided by Hann, Rock, King, and Drewnowski (2001) who found that higher HEI scores were associated with biomarkers of dietary intake.

Data analyses

Given the longitudinal nature of the data, latent growth curve modeling as recommended by Ram and Grimm (2007), was conducted using M-Plus (M-Plus 6.1, Muthén & Muthén, 2006). Latent growth curve (LGC) modeling allows us to explain intra-individual change over time by modeling specific trajectories of how this change occurred. The main advantage in LGC modeling is that any number of differing patterns of change can be modeled (Ram & Grimm, 2007). Then, between-subject differences in change over time can be examined.

First, changes during the intervention were modeled in each of the four predictors (benefits, barriers, self-efficacy, and family support, models 01e04). The criteria for inclusion in later testing were that the variable had to show significant change over the intervention and a significant treatment group effect. Next, models were constructed to examine intra-individual change in healthy eating behavior in response to the intervention in both groups (Models 05-09). These analyses were followed by models to examine between-subject differences in the outcome, in change over time, and what determinants influence those changes at each phase of the intervention and follow-up (Models 10-13). The final analyses aimed to explain whether significant group effects (coded tailored group = 1 and standard = 2) in changes over time in eating behavior were mediated by changes in the predictors (Preacher & Hayes, 2004). The significance level for the study was set at .05. The amount of missing data in the Healthy Eating Index at

each time point (0, 3, 6, 9, 12, 18, and 24 months) was: 0%, 2%, 5%, 7%, 5%, 6%, and 5%, respectively. M-Plus uses FIML (full information maximum likelihood) to impute missing data.

Results

Women in both groups were, on average, 58 years old (SD = 5.5; Range = 50-69), Caucasian, married, employed outside the home, had attended some college or were college graduates, had an average BMI of 30 which did not change over time, and reported annual incomes >\$20,000 (Walker et al., 2009). Summary scores for each predictor variable were calculated by group (Table 1).

Table 1
Summary scores of predictor variables by intervention group.

	Baseline	3 month	6 month	9 month	12 month				
	Mean (SD)								
Family support (Range 1-5)									
Tailored group	1.75 (.80)	2.18 (.96)	2.11 (.91)	2.16 (.98)	2.05 (.97)				
Standard group	1.67 (.80)	1.90 (.82)	1.83 (.80)	1.85 (.92)	1.82 (.86)				
Barriers (Range 1-4)									
Tailored group	2.19 (.49)	2.04 (.44)	2.03 (.47)	2.01 (.46)	2.01 (.49)				
Standard group	2.25 (.45)	2.18 (.46)	2.12 (.37)	2.15 (.37)	2.16 (.43)				
Benefits (Range 1-4)									
Tailored group	3.28 (.42)	3.28 (.41)	3.34 (.40)	3.34 (.43)	3.38 (.45)				
Standard group	3.30 (.42)	3.32 (.43)	3.24 (.41)	3.24 (.46)	3.19 (.41)				
Self-efficacy (Range 1-5)									
Tailored group	3.57 (.79)	3.71 (.77)	3.76 (.79)	3.73 (.82)	3.75 (.83)				
Standard group	3.48 (.77)	3.52 (.72)	3.58 (.79)	3.62 (.85)	3.60 (.83)				
					1				

Change over time in predictors

A separate latent growth curve model was created for each predictor testing for differences during the intervention. Model fit statistics are provided in Table 2. For self-efficacy and benefits for healthy eating, no significant change over time or any significant differences between groups were observed. For family support, a

significant increase over the first year of the intervention was observed (b = .160, β = .599, z = 2.82, p < .05). At the end of the intervention, the tailored newsletter group reported significantly more family support than the standard newsletter group (b = -.289, β = -.366, z = 2.40, p < .05). For perceived barriers, a significant decrease was observed during the intervention (b = -.101, β = -.455, z = 3.34, p < .05) and the tailored group reported significantly fewer barriers than the standard group (b = .140, β = .369, z = 2.42, p < .05) at the end of the intervention. Based on these results, only family support and barriers were used in later testing of changes over time in healthy eating.

Table 2 Summary of model fit statistics.

Model details		c² (df)	CFI	RMSEA	SRMR				
Changes in the predictors									
01	Self-efficacy for healthy eating	14.60 (13)	.99	.02	.03				
02	Benefits	6.68 (13)	1.00	.00	.04				
03	Family support	17.88 (13)	.99	.05	.06				
04	Perceived barriers	12.89 (13)	.99	.04	.09				
Intra-individual models									
05	Unconditional model	1100.70 (21)*	N/A	.48	.51				
06	Linear change	106.47 (23)*	.92	.13	.11				
07	Revised linear model	30.55 (14)*	.98	.07	.07				
80	Curvilinear change	82.59 (19)*	.94	.12	.06				
09	Revised curvilinear model	18.70 (10)	.99	.06	.03				
Inter-individual models									
10	Main effects of newsletter	19.28 (13)	.99	.05	.03				
11	Barriers for healthy eating	80.73 (44)*	.98	.06	.06				
12	Family support	147.84 (110)*	.98	.04	.06				
13	Mediational model	157.02 (124)*	.99	.03	.06				

^{*}p < .05.

Note. N/A = Not applicable.

CFI = Comparative fit index.

RMSEA = Root mean square error of approximation.

SRMR = Standardized root mean square residual.

Intra-individual change

Next, change over time in healthy eating behavior was examined (Fig. 1). We tested an unconditional model (Model 05, Table 2) which served as a basis for comparison in the model building process. Model 06 examined simple linear change over time (whether healthy eating improved over all 7 time points; Fig. 2). Model 07 tested a revised linear model to separate change during the intervention from

change during the maintenance period. The linear models proved a poor fit to the data. Model 08 examined curvilinear change over time (7 time points) and Model 09 separated change during the intervention from change during follow- up. This last model represented the best statistical and theoretical fit to the data.

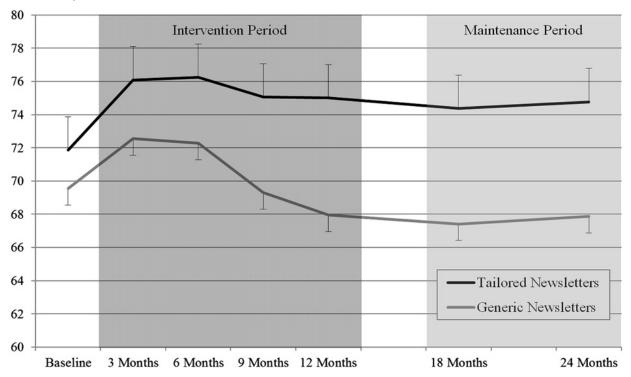


Fig. 1. Raw mean values of the healthy eating index by group during the intervention (3-12 months) and maintenance (18 and 24 months) periods (with standard errors). The darker shading represents the intervention period and the lighter shading represents the maintenance period.

Inter-individual differences

Next, we examined whether groups differed in healthy eating change as a result of the intervention, at the end of the intervention, and during follow-up. Model 10 reflects the addition of treatment group to the model (Fig. 3). The tailored newsletter group increased healthy eating more than the standard newsletter group (b = -6.661, $\beta = -.452$, z(1-tailed) = 1.73, p < .05). Statistically significant differences were observed at the end of the intervention (b = -7.199, $\beta = -.705$, z = 4.93, p < .05) and in maintenance (b = -7.245, $\beta = -.721$, z = 5.12, p < .05). No significant group difference was observed in the curvilinear effects (b = -1.832, $\beta = -.114$, z = .482, p > .05).

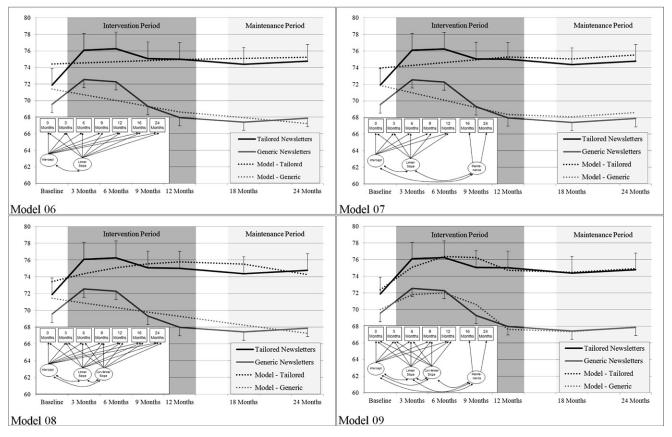
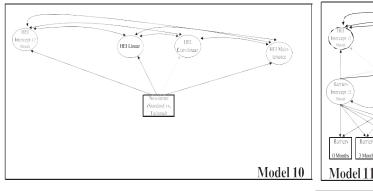


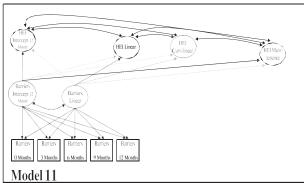
Fig. 2. Fitted lines plotted against the mean scores of the healthy eating index (HEI) at each time point. The solid lines denote raw mean scores. The dotted lines denote fitted scores. Model 06 examined simple linear change over time using all 7 time points; model 07 was a revised linear model separating change during the intervention from change during maintenance. Model 08 examined curvilinear change over time using all 7 time points; model 09 was a revised curvilinear model separating change during the intervention from change during maintenance (best statistical and theoretical fit to the data). HEI Intercept was rescaled to represent healthy eating at the end of the intervention. HEI linear denotes the intra-individual linear change over time. HEI curvilinear denotes the intra-individual curvilinear change over time. HEI maintenance denotes healthy eating index scores during the maintenance period.

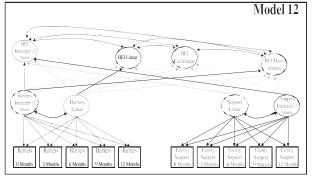
Model 11 reflects the addition of barriers to the model. Decreases in barriers during the intervention were significantly related to increases in healthy eating over this same time frame (b = -47.438, β = -.643, z(1-tailed) = 1.66, p < .05). Lower barriers at the end of the intervention were related to better healthy eating scores at the end of the intervention (b = -7.286, β = -.269, z = 3.08, p < .05) and resulted in higher HEI scores during the maintenance period (b = -10.775, β = -.410, z = 5.16, p < .05).

Model 12 reflects the addition of family support. Healthy eating at the end of the intervention was positively associated with family support at the end of the intervention (b = 1.500, $\beta = .116$, z = 2.39, p < .05). No significant effects were observed for

change over time in family support on changes in the HEI outcomes. Separate post-hoc analyses revealed a significant association between healthy eating during the second year and family support at the same time (b = 2.086, $\beta = .211$, z = 2.45, p < .05). Lastly, analyses for latent moderation were done to test whether the effect of barriers on healthy eating differed as a function of differing levels of family support. No significant interactions were observed suggesting that the effects in the current analyses were independent of each other.







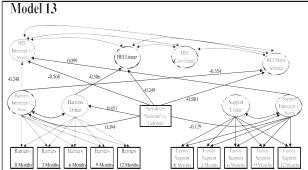


Fig. 3. Model building process for explaining change over time in rural midlife and older women's healthy eating behavior. Note. Standardized beta coefficients are provided for the significant values in models 10 & 11; models 12 & 13 were included to depict model construction. In model 13, variance explained is by each latent construct. HEI = healthy eating index (the indicators of healthy eating for the 7 assessments have been removed for clarity). Intercept was rescaled to represent healthy eating at the end of the intervention for HEI, barriers, and family support. Linear denotes the intra-individual linear change over time. Curvilinear denotes the intra-individual curvilinear change over time. HEI maintenance: denotes the healthy eating index score during the maintenance period.

Tests for mediation

The final set of analyses examined whether the observed group differences in healthy eating were mediated by the group effects on family support and perceived barriers (Model 13). Part of the treatment group effect on healthy eating at the end of

the intervention was through a decrease in barriers (b = -1.003, $\beta = -.098$, z = 2.07, p < .05) and an increase in family support (b = -.420, $\beta = -.041$, z(1-tailed) = 1.74, p < .05) reflecting a 19.43% drop in the standardized coefficient. Finally, part of the group effect on healthy eating during maintenance was also through a decrease in barriers by the end of the intervention (b = 1.384, $\beta = -.139$, z = 2.23, p < .05) reflecting a 19.56% drop in the standardized coefficient.

Discussion

Healthy eating as measured by the HEI increased in both groups with the majority of change occurring in the first six months. The tailored newsletter group increased their healthy eating more than the standard newsletter group during and at the end of the intervention, and during follow-up. It is worth noting that no significant group difference was observed in the curvilinear effects suggesting that both groups decreased similarly after making initial gains. Although healthy eating scores dropped for both groups after 6 months, the tailored newsletter group maintained their healthy eating behavior at a significantly higher level than the standard newsletter group. This latter group dropped below their baseline score (M=69.5) at month 9 and remained below baseline throughout the 24 month follow-up. In the evaluation of the study, women mentioned that they learned how to be more accurate at completing the food frequency questionnaire during the study, which may explain why scores dropped in both groups after 6 months. Future studies could potentially train participants in accurate reporting of eating habits before starting an intervention.

Using a composite measure of eating behavior may have captured the robust differences between groups in eating behavior better than single food categories (i.e., fruits & vegetables). In several studies, the mean HEI for women of varying ages was as low as 42.8 for low-income women, ages 19e50 (Duffy, Zizza, Jacoby, & Tayie, 2009); but typically was in the 60's for women ages 51 and older (Boynton et al., 2007; Grimstvedt, Woolf, Milliron, & Manore, 2009; Savoca et al., 2009; Tande, Magel, & Strand, 2009). This would place most midlife to older women in a fair/good classification for healthy eating according to this index. The HEI results for the women in our study were slightly above earlier results.

It appears that providing rural women with tailored messages had a greater influence on patterns of change in healthy eating behavior than the standard newsletters. From prior studies, tailored interventions typically have longer lasting and more powerful effects than standard newsletters (Noar et al., 2007). In this study, the tailored group was able to sustain their eating behavior changes better than the standard group in part because of the individualized information they received. Although prior literature has found that rural dwellers do not use preventive services to the same extent as their counterparts (Bove & Olson, 2006; McCall-Hosenfeld & Weisman, 2011; Paluck, Allerdings, Kealy, & Dorgan, 2006; Pullen et al., 2001), it is likely that the individualized counseling about lifestyle behavior delivered via an easily accessible method that could be reviewed at their convenience were key factors in bringing about the change.

Three significant effects were observed for the influence of family support. First, family support at the end of the intervention was positively associated with healthy eating at this same time point. Second, family support during the follow-up year was also related to healthy eating during follow-up. Third, during the intervention, family support increased more in the tailored news- letter group than the standard newsletter group. Women in this group were able to garner more family support as a direct result of the tailored intervention which, in turn, was related to healthier eating for them at the end of the intervention. The six newsletters that were devoted to interpersonal support covered such topics as how to ask for support for healthy eating, accepting help that is provided, ideas about who to ask for support, and so forth. In earlier studies, more family support was related to higher levels of healthy eating in women (Chang, Baumann, Nitzke, & Brown, 2005; Luszczynska & Cieslak, 2009; Steptoe, Perkins-Porras, Rink, Hilton, & Cappuccio, 2004). The current findings support the cultural belief that rural women rely on their families for assistance with health-promoting behaviors (Long & Weinert, 1989; Paluck et al., 2006). In contrast, two studies found that family support for nutrition was related to healthier eating but this effect was, in large part, indirect through other predictors (Anderson, Winett, & Wojcik, 2007; Luszczynska & Cieslak, 2009). Similarly, Hermstad et al. (2010) did not find a relationship between family support for healthy eating and dietary fat intake in their study of rural men and women. The role

of family support for healthy eating warrants further study.

It was surprising that self-efficacy for healthy eating did not differ between the treatment groups and was not related to healthy eating behavior in the current study. Although earlier investigators also have not found a significant relationship between these variables (Brug, de Vet, de Nooijer, & Verplanken, 2006; Kellar & Abraham, 2005), self-efficacy, more than any other determinant, has been related to healthy eating behavior (Anderson et al., 2007; Folta et al., 2009; Fuemmeler et al., 2006; Luszczynska & Cieslak, 2009). In this study, self-efficacy for healthy eating was the strategy given the most emphasis, discussed in 16 of 18 newsletters. Despite this emphasis, the mean scores for self-efficacy were fairly high in both groups and remained stable over time. This suggests that there may have been a ceiling effect for self-efficacy such that both groups already had a "threshold" level of self-efficacy and other psychosocial factors were more salient. Alternatively, it may be that increasing the women's confidence for healthy eating may have required a stronger, more intense intervention than a mailed newsletter (Lin et al., 2010).

Several effects were observed for the influence of perceived barriers. First, decreases in barriers over the intervention were related to increases in healthy eating over this same time. Second, lower barriers at the end of the intervention were related to better healthy eating scores at the end of the intervention. Third, lower barriers at the end of the intervention resulted in higher healthy eating scores during maintenance. In addition, participants in the tailored newsletter group perceived fewer barriers to healthy eating than those in the standard newsletter group at the end of the intervention and during the maintenance phase, which resulted in healthier eating. It appears that the tailored newsletters were effective in counteracting many of the barriers that women perceived during the intervention (inconvenience, expense, time required for shopping and preparing healthy meals, and giving up preferred foods). Similar to this study, several earlier studies found that barriers were negatively related to healthy eating behavior (Brug et al., 2006; Kaiser et al., 2010). In focus group participants, Rolnick et al. (2009) reported that barriers to fruit and vegetable consumption were lack of time, expense, and fruit and vegetable availability.

Perceived benefits of healthy eating did not differ between the treatment groups or relate to healthy eating behavior in this study. Benefits of healthy eating have not shown as consistent an effect on eating behavior in prior research as other determinants (Shaikh, Yaroch, Nebeling, Ming-Chin, & Resnicow, 2008). Compared to self-efficacy or social support, it may be that the benefits of eating healthy foods are too distant in time to have a strong influence on the women's eating behavior as suggested by Hagler et al. (2007).

This study has several limitations. The composite outcome used in the current analysis (HEI instead of separate food categories), made this analysis difficult to compare with the earlier published results from this study. A further limitation was the varying time frame for measuring the food frequency questionnaire from every three months during the intervention, to every six months during maintenance, and the past year at baseline. Although the women's BMI did not change over time, future studies should examine the effects of changes in eating behavior on BMI. Data were based on self-report and women may not have recalled eating behaviors accurately. Although we used cluster randomization to assign the two rural counties to intervention groups, we were unable to take into account the effect of clustering because we only had one cluster (county) per group. Future studies need to account for effects of clustering. Because we were not able to collect baseline data on nonparticipants, it was unknown how they differed from participants. Women generally lived within 50 miles of the rural research office and thus had fewer barriers to travel compared to other rural women. Thus, the generalizability of the findings is limited to midlife/older women who reside in rural areas similar to this study. More research is needed to clarify the effects of selfefficacy and benefits on healthy eating behavior in rural women as these were not significant in this study.

In summary, the tailored newsletter intervention was successful in improving participants' healthy eating behavior compared to the standard newsletter intervention during and at the end of the intervention, and during maintenance. Perceived barriers had the strongest impact on healthy eating behavior during and at the end of the intervention and the maintenance period. Participants in the tailored newsletter group perceived more family support and fewer barriers for healthy eating at the end of the

intervention than those in the standard newsletter group. Thus, both family support and barriers should be a major focus of interventions to improve healthy eating behavior in rural midlife and older women.

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